

# Scalable Bayesian seismic wavelet estimation

Guillermina Senn<sup>1</sup>

Joint work with Matt Walker<sup>2</sup>, Håkon Tjelmeland<sup>1</sup>, and Andrew Holbrook<sup>3</sup>

<sup>1</sup>Norwegian University of Science and Technology (NTNU)

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SINTEF Industri, Trondheim

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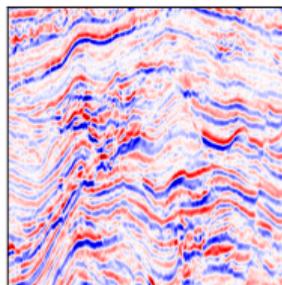
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# Summary

## 1. Seismic data



- Reflectivity  $\mathbf{c}$
- Seismic wavelet  $\mathbf{w}$
- Observational noise  $\mathbf{e}$
- Seismic data  $\mathbf{d}$

## 2. Model

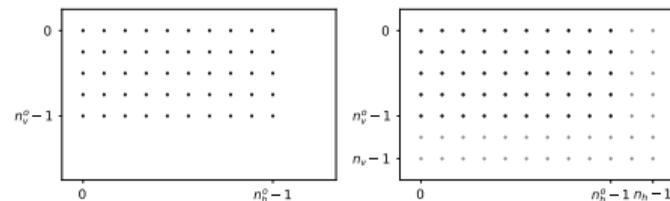
1D convolutional model:

$$\mathbf{d}_j = \mathbf{c}_j \star \mathbf{w} + \mathbf{e}_j \quad \text{for column } j.$$

Probabilistic model:

- Gaussian reflectivity prior.
- Gaussian wavelet prior.
- Gaussian likelihood.
- Inverse-Gamma variances.

## 3. Scalability



## 4. Estimation

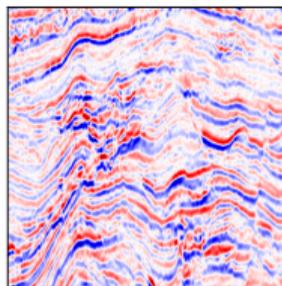
Sample posterior distribution

$$p(\mathbf{c}, \mathbf{w}, \sigma_c^2, \sigma_w^2, \sigma_d^2 \mid \mathbf{d})$$

with Markov chain Monte Carlo (MCMC).

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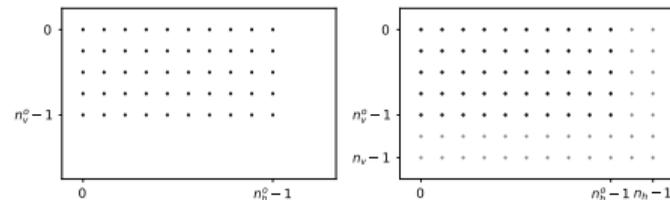
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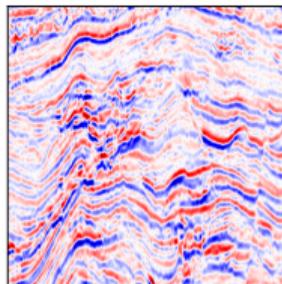
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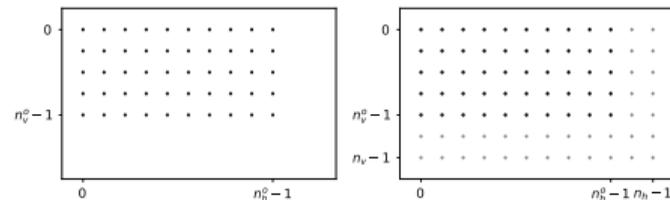
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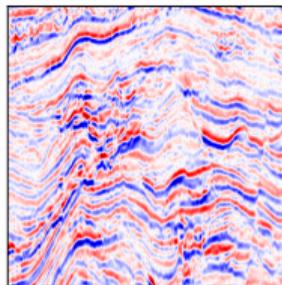
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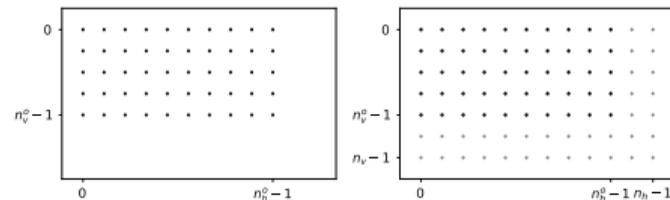
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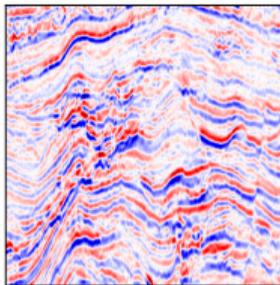
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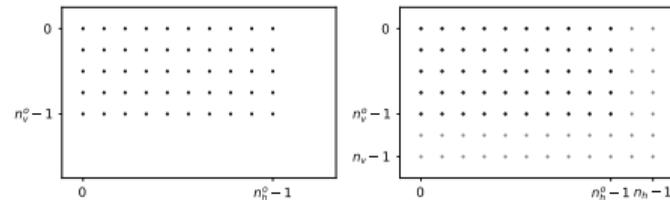
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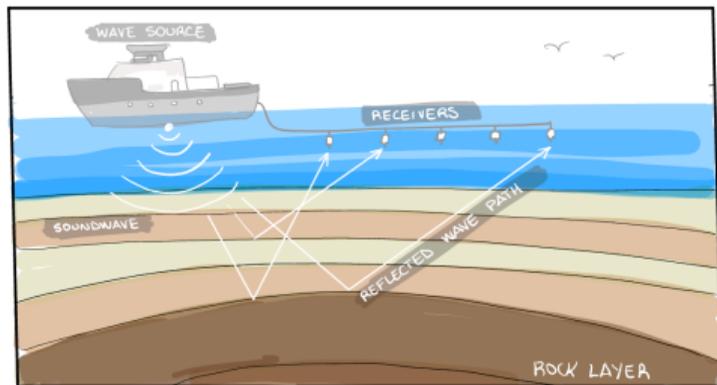
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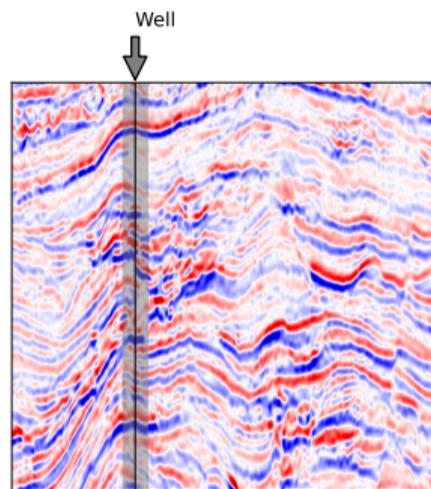
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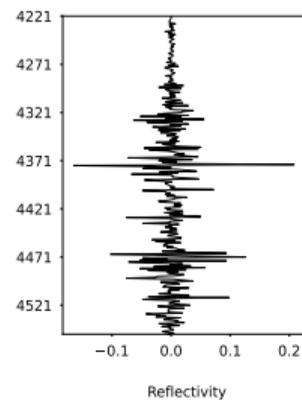
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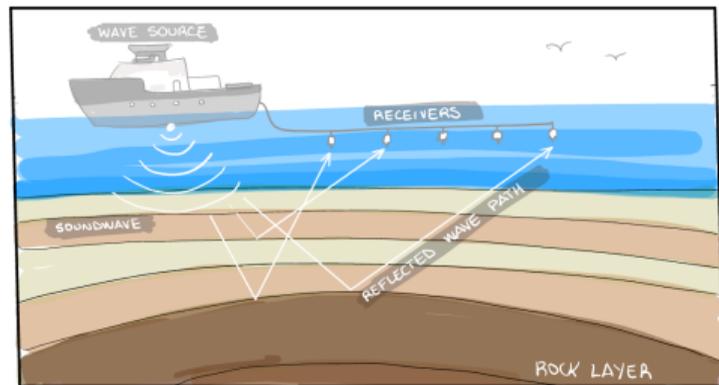


(b) Seismic data  $D$

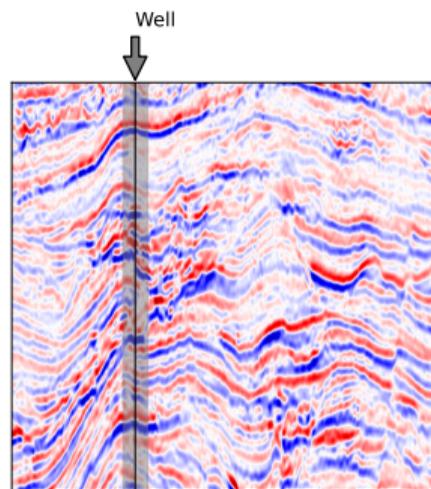


(c) Well log data  
 $c_{\text{well}}$ .

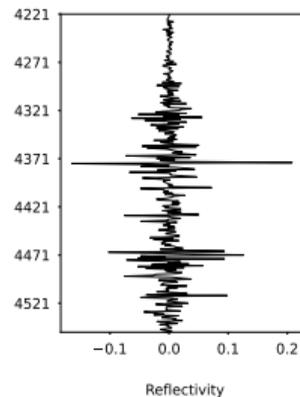
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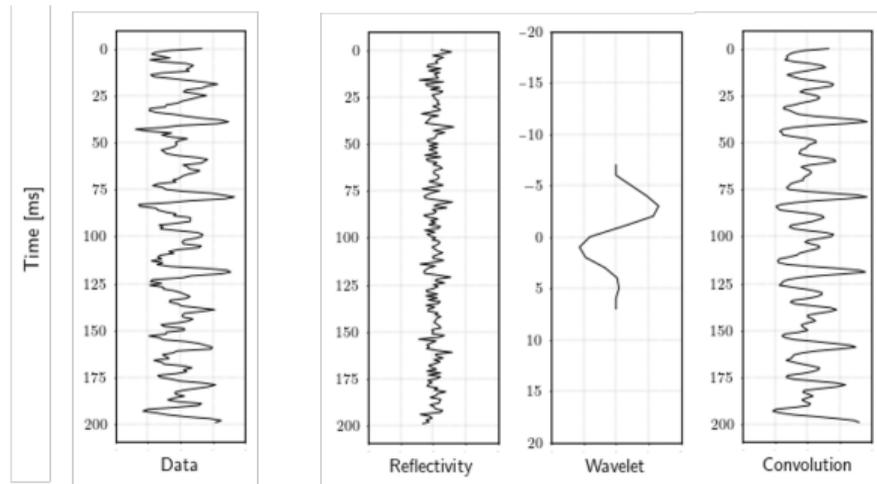


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## Model assumptions



1D convolutional model for column  $j$ :

$$\mathbf{d}_j = \mathbf{w} \star \mathbf{c}_j + \mathbf{e}_j, \quad j = 1, \dots, m$$

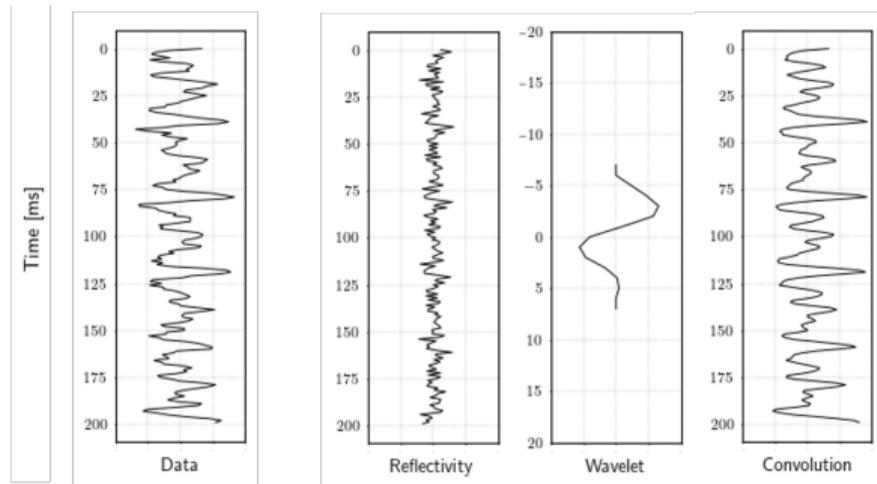
- Denote the data, reflectivity, and noise for all columns on the image by  $\mathbf{D}$ ,  $\mathbf{C}$ , and  $\mathbf{E}$ .
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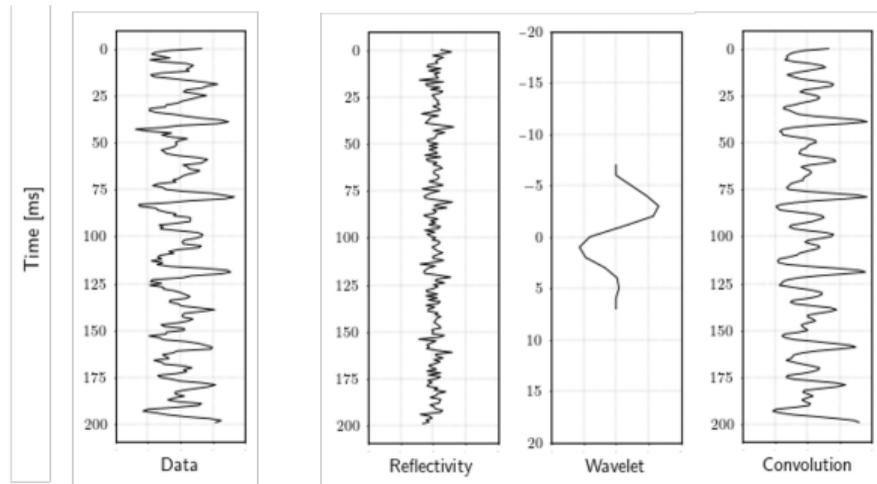
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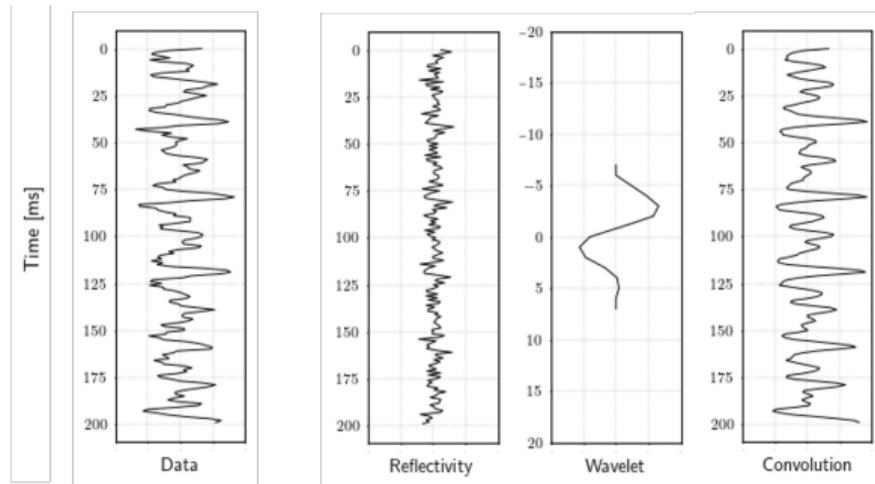
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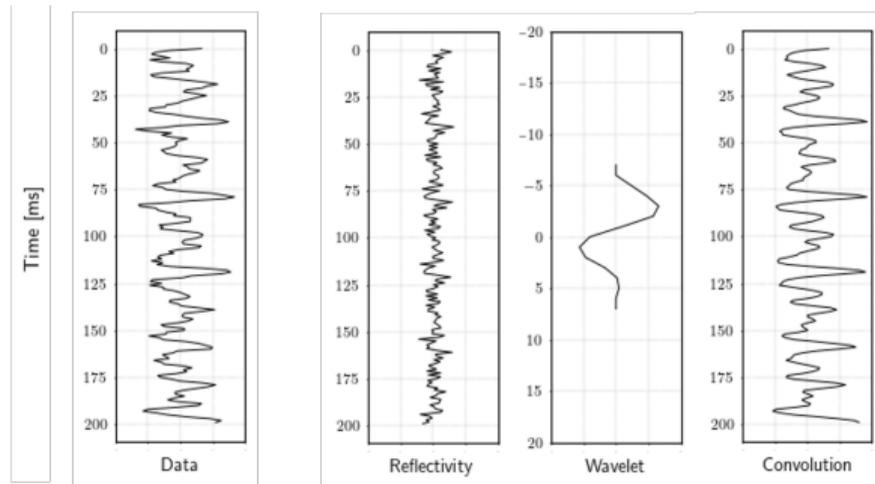
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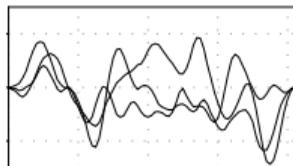
## The probabilistic model (1)

- Represent the wavelet in the time domain as the sequence with  $k$  elements

$$\mathbf{w} = \{w_1, \dots, w_k\}.$$

- Define the  $2 \times k$  constraint matrix  $\mathbf{A}_w$  that forces the endpoints of  $\mathbf{w}$  to be zero.
- Model the constrained sequence as a Gaussian process:

$$\mathbf{w} | \sigma_w^2 \sim N_k(0, \sigma_w^2 \mathbf{R}_w), \quad \mathbf{w}^* = \mathbf{w} | \mathbf{A}_w \mathbf{w} = \mathbf{0}, \quad \rightarrow \mathbf{w}^* | \sigma_w^2 \sim N_k(0, \sigma_w^2 \mathbf{R}_w^*)$$



Samples from (constrained) wavelet prior.

- Similarly, model the constrained reflectivity vector as a Gaussian field:

$$\mathbf{c} | \sigma_c^2 \sim N_{nm}(0, \sigma_c^2 \mathbf{R}_c), \quad \mathbf{c}^* = \mathbf{c} | \mathbf{A}_c \mathbf{c} = \mathbf{c}_{\text{well}} \quad \rightarrow \mathbf{c}^* | \sigma_c^2 \sim N_{nm}(\boldsymbol{\mu}_c^*, \sigma_c^2 \mathbf{R}_c^*)$$

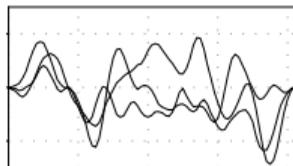
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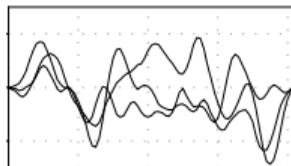
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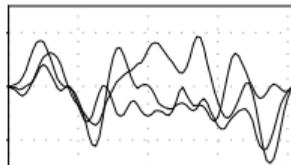
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## The probabilistic model (2)

- Assume Gaussian observational noise and obtain the Gaussian likelihood

$$\mathbf{d}|\mathbf{c}, \mathbf{w}, \sigma_c^2, \sigma_w^2, \zeta \sim N_{nm}(\mathbf{W}\mathbf{c}, \sigma_d^2 \mathbf{R}_d), \quad \sigma_d^2 \propto \sigma_c^2 \sigma_w^2 \zeta, \quad \zeta^{-1} = \text{SNR}$$

- Model the spatial correlations as the product of the correlations in each direction (separability):

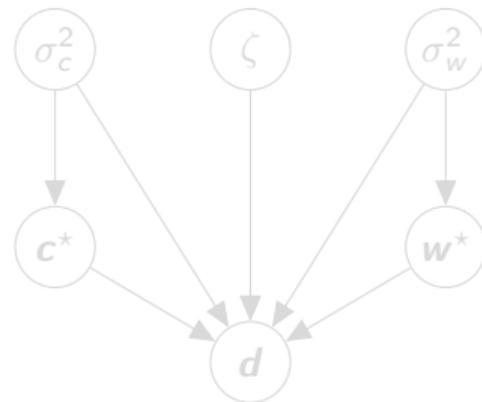
$$\mathbf{R}_d = \mathbf{R}_{d,h} \otimes \mathbf{R}_{d,v}, \quad \mathbf{R}_c = \mathbf{R}_{c,h} \otimes \mathbf{R}_{c,v}.$$

- Assign inverse-Gamma hyperpriors to the marginal variance parameters

$$\sigma_c^2 \sim IG(\alpha_c, \beta_c)$$

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The stochastic model represented by a DAG.

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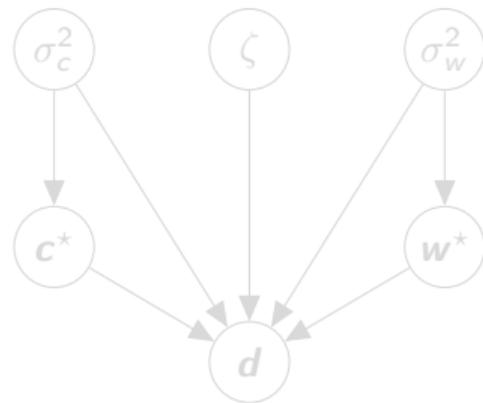
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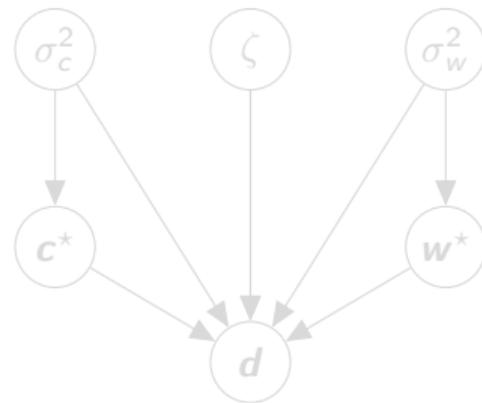
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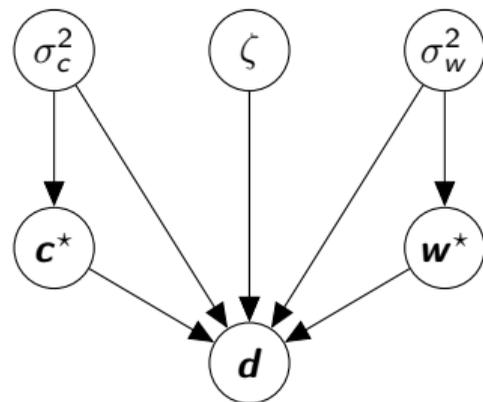
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## Posterior distribution of the model

- Denote: data  $\mathbf{y} = (\mathbf{d}, \mathbf{c}_{well})$  and unknown parameters  $\boldsymbol{\theta} = (\mathbf{w}^*, \mathbf{c}^*, \sigma_w^2, \sigma_c^2, \zeta)$ .
- Use Bayes' rule

$$p(\boldsymbol{\theta}|\mathbf{y}) = \frac{p(\boldsymbol{\theta}, \mathbf{y})}{p(\mathbf{y})} \propto p(\mathbf{y}, \boldsymbol{\theta})$$

to find the joint distribution

$$p(\mathbf{c}^*, \mathbf{w}^*, \sigma_c^2, \sigma_w^2, \zeta|\mathbf{d}) = p(\mathbf{d}|\mathbf{c}^*, \mathbf{w}^*, \sigma_c^2, \sigma_w^2, \zeta)p(\mathbf{c}^*|\sigma_c^2)p(\mathbf{w}^*|\sigma_w^2)p(\sigma_c^2)p(\sigma_w^2)p(\zeta).$$

- Conjugate distributions in the model (Gaussian-Gaussian and Gaussian-inverse-Gamma) let us write a Gibbs sampler.

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- Conjugate distributions in the model (Gaussian-Gaussian and Gaussian-inverse-Gamma) let us write a Gibbs sampler.

## Posterior distribution of the model

- Denote: data  $\mathbf{y} = (\mathbf{d}, \mathbf{c}_{well})$  and unknown parameters  $\boldsymbol{\theta} = (\mathbf{w}^*, \mathbf{c}^*, \sigma_w^2, \sigma_c^2, \zeta)$ .
- Use Bayes' rule

$$p(\boldsymbol{\theta}|\mathbf{y}) = \frac{p(\boldsymbol{\theta}, \mathbf{y})}{p(\mathbf{y})} \propto p(\mathbf{y}, \boldsymbol{\theta})$$

to find the joint distribution

$$p(\mathbf{c}^*, \mathbf{w}^*, \sigma_c^2, \sigma_w^2, \zeta|\mathbf{d}) = p(\mathbf{d}|\mathbf{c}^*, \mathbf{w}^*, \sigma_c^2, \sigma_w^2, \zeta)p(\mathbf{c}^*|\sigma_c^2)p(\mathbf{w}^*|\sigma_w^2)p(\sigma_c^2)p(\sigma_w^2)p(\zeta).$$

- Conjugate distributions in the model (Gaussian-Gaussian and Gaussian-inverse-Gamma) let us write a Gibbs sampler.

# Gibbs sampler for joint reflectivity and wavelet estimation

---

## Algorithm Gibbs sampler

---

**Input:** Seismic data  $\mathbf{d}$ , well log  $\mathbf{c}_{\text{well}}$ , priors, initial values

**Output:** Samples from posterior  $p(\mathbf{c}^*, \mathbf{w}^*, \sigma_c^2, \sigma_w^2, \zeta \mid \mathbf{d})$

Initialize  $\mathbf{c}^{(0)}, \mathbf{w}^{(0)}, (\sigma_c^2)^{(0)}, (\sigma_w^2)^{(0)}, \zeta^{(0)}$

**for**  $t = 1, 2, \dots, T$  **do**

**Update reflectivity:** Sample  $\mathbf{c}^{(t)} \sim p(\mathbf{c}^* \mid \mathbf{w}^{(t-1)}, (\sigma_c^2)^{(t-1)}, (\sigma_w^2)^{(t-1)}, \zeta^{(t-1)}, \mathbf{d})$

**Update wavelet:** Sample  $\mathbf{w}^{(t)} \sim p(\mathbf{w}^* \mid \mathbf{c}^{(t)}, (\sigma_c^2)^{(t-1)}, (\sigma_w^2)^{(t-1)}, \zeta^{(t-1)}, \mathbf{d})$

**Update reflectivity variance:** Sample  $(\sigma_c^2)^{(t)} \sim p(\sigma_c^2 \mid \mathbf{c}^{(t)}, \mathbf{w}^{(t)}, (\sigma_w^2)^{(t-1)}, \zeta^{(t-1)}, \mathbf{d})$

**Update wavelet variance:** Sample  $(\sigma_w^2)^{(t)} \sim p(\sigma_w^2 \mid \mathbf{c}^{(t)}, \mathbf{w}^{(t)}, (\sigma_c^2)^{(t)}, \zeta^{(t-1)}, \mathbf{d})$

**Update SNR:** Sample  $\zeta^{(t)} \sim p(\zeta \mid \mathbf{c}^{(t)}, \mathbf{w}^{(t)}, (\sigma_c^2)^{(t)}, (\sigma_w^2)^{(t)}, \mathbf{d})$

**Return:**  $\{\mathbf{c}^{(t)}, \mathbf{w}^{(t)}, (\sigma_c^2)^{(t)}, (\sigma_w^2)^{(t)}, \zeta^{(t)}\}_{t=B+1}^T$  after burn-in  $B$

---

The problem is that the full conditionals  $p(\theta_j \mid \theta_{-j}, \mathbf{d})$  involve operations with computational complexity  $\mathcal{O}(l^3)$ ,  $l = nm$ .

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## A possible solution

- Matrix multiplication, inversion, and decomposition for arbitrary  $ss$  matrices scale  $\sim$  cubic with  $s$ .
- For matrices with special structures we have efficient multiplication/inversion algorithms.
- For example, take the circulant matrix

$$\underline{R} = \text{circ}(r) = \begin{pmatrix} r_0 & r_1 & r_2 & \cdots & r_{s-1} \\ r_{s-1} & r_0 & r_1 & \cdots & r_{s-2} \\ r_{s-2} & r_{s-1} & r_0 & \cdots & r_{s-3} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ r_1 & r_2 & r_3 & \cdots & r_0 \end{pmatrix}, \quad r = (r_0, \dots, r_{s-1})^T.$$

- We can represent it by its base  $r$ .
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Can we make our correlation matrices circulant?

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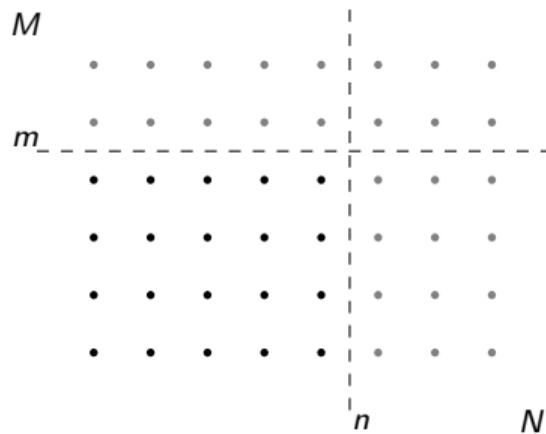
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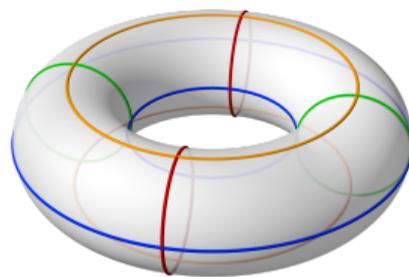
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## Scalability



(a) The extended lattice.



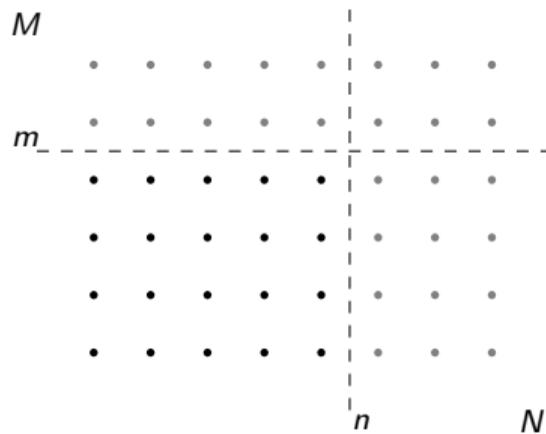
(b) The cyclic lattice.

- Separability on the cyclic lattice gives circulant or block-circulant with circulant blocks correlation matrices:

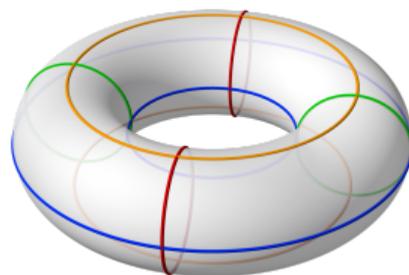
$$\underline{\underline{R}}_{w,v}, \quad \underline{\underline{R}}_c = \underline{\underline{R}}_{c,h} \otimes \underline{\underline{R}}_{c,v}, \quad \underline{\underline{R}}_d = \underline{\underline{R}}_{d,h} \otimes \underline{\underline{R}}_{d,v}$$

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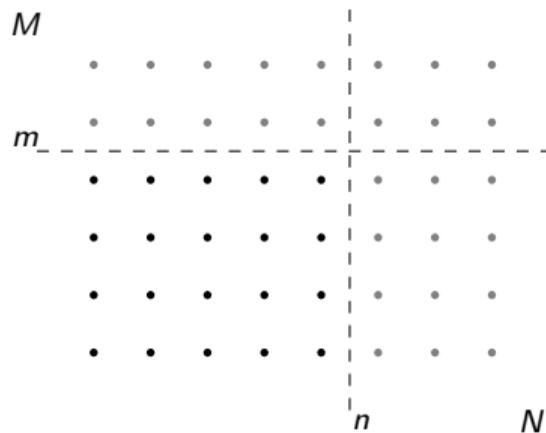
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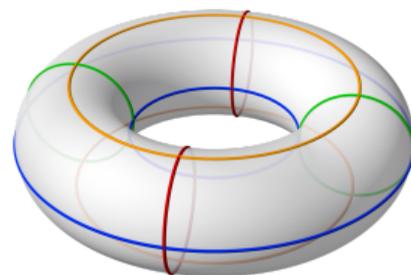
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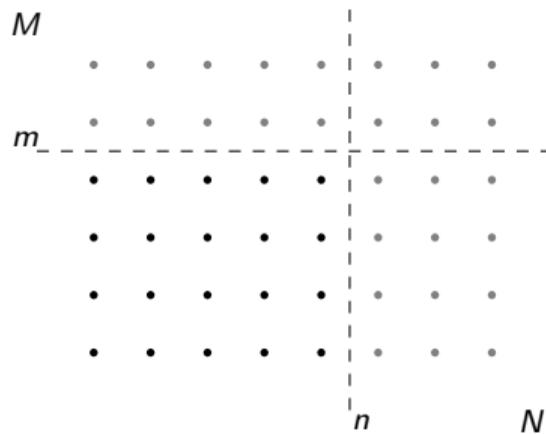
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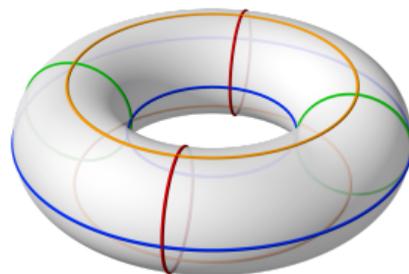
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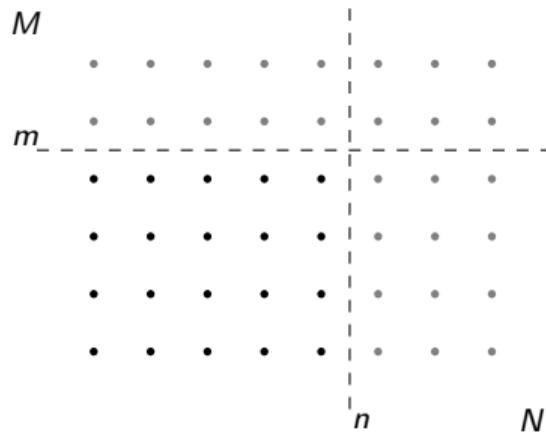
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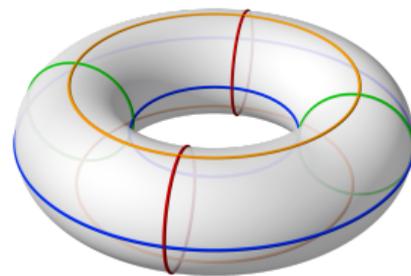
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# Gibbs sampling on the cyclic lattice

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## Algorithm Gibbs sampler on the cyclic lattice

---

**Input:** ...

**Output:** Samples from posterior  $p(\mathbf{c}^*, \mathbf{w}^*, \sigma_c^2, \sigma_w^2, \zeta, \mathbf{d}_{\text{aux}} \mid \mathbf{d})$

Initialize ...

**for**  $t = 1, 2, \dots, T$  **do**

**Update reflectivity**

**Update wavelet**

**Update reflectivity variance**

**Update wavelet variance**

**Update SNR**

**Update auxiliary seismic data  $\mathbf{d}_{\text{aux}}$**

**Return:**  $\{\mathbf{c}^{(t)}, \mathbf{w}^{(t)}, (\sigma_c^2)^{(t)}, (\sigma_w^2)^{(t)}, \zeta^{(t)}, \mathbf{d}_{\text{aux}}^{(t)}\}_{t=B+1}^T$  after burn-in  $B$

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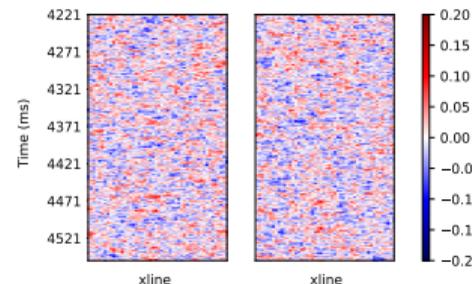
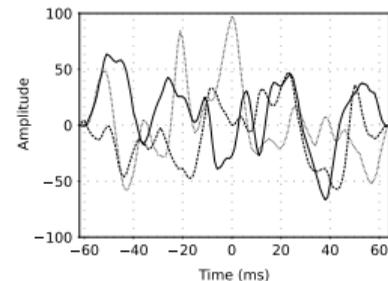
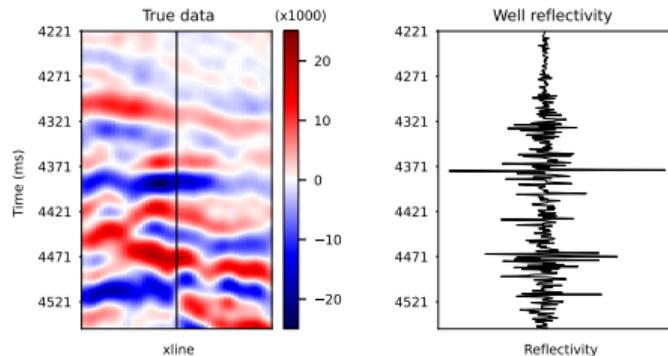
Some advantages:

- Store only bases of  $\underline{\mathbf{R}}_{\cdot, h}$ ,  $\underline{\mathbf{R}}_{\cdot, v}$ ,  $\underline{\mathbf{R}}_{\cdot}$ .
- Use DFT to compute full conditional parameters and sample from Gaussians in  $\mathcal{O}(S \log(S))$ ,  $S = NM$ .

# Real application: Gas reservoir in offshore Egypt

We extract a  $330 \times 50$ s window from the AVO data for a fixed inline, centered around the well.

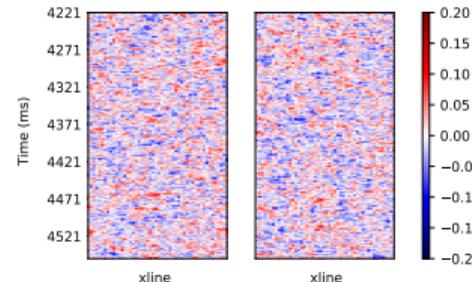
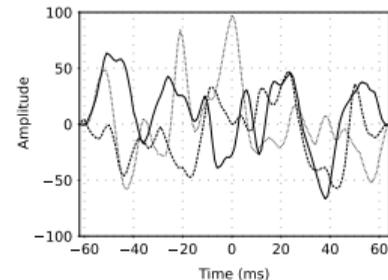
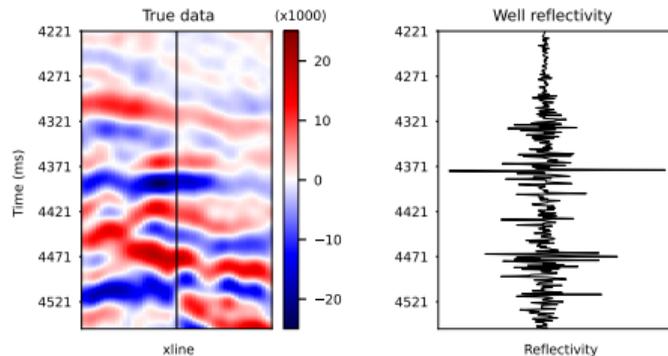
We choose high correlation in the wavelet prior and moderate correlation in the reflectivity prior. Some prior samples:



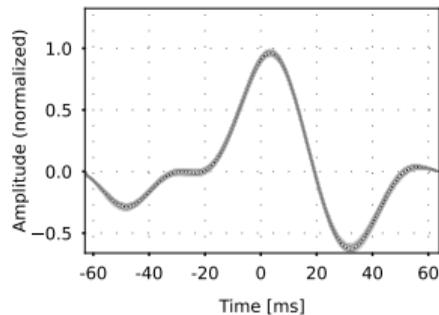
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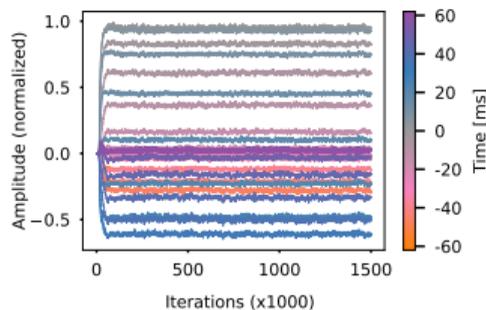
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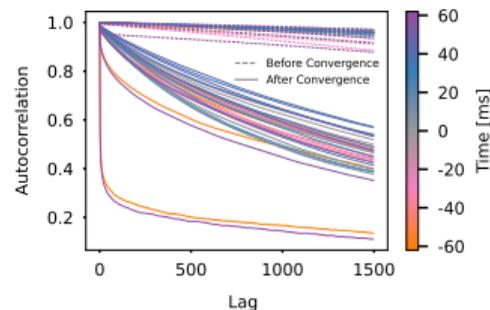
# Wavelet posterior exploration



(a) Posterior samples



(b) Traceplots

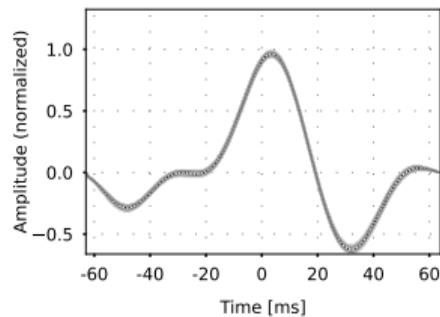


(c) Sample autocorrelation

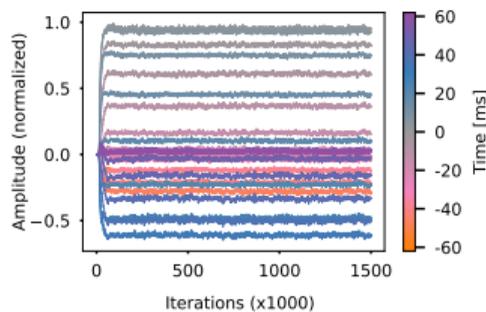
MCMC sampling results for wavelet posterior distribution.

Question: Can we sample more efficiently from the same posterior distribution?

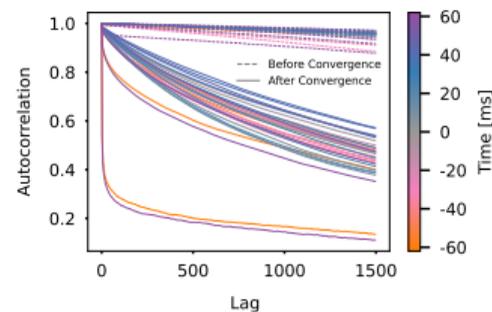
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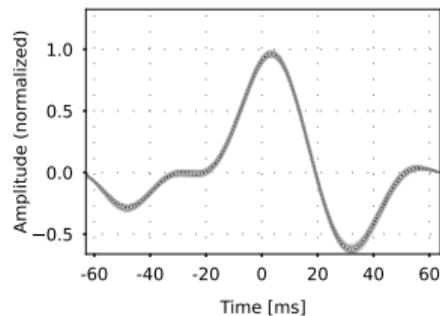


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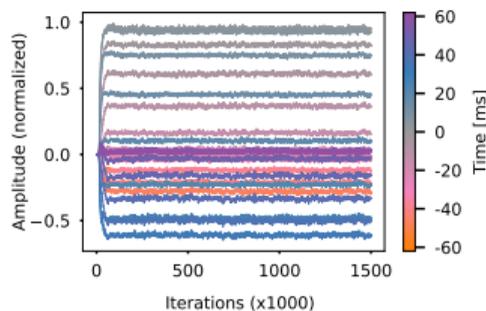
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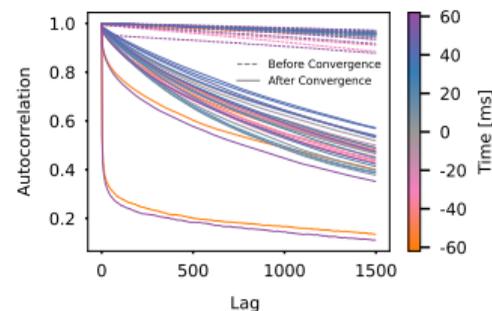
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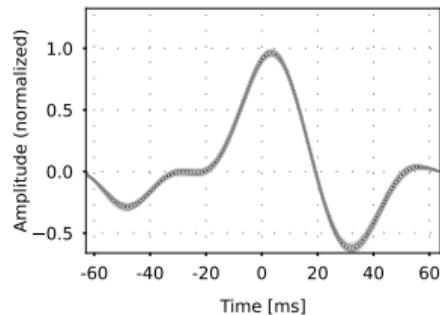


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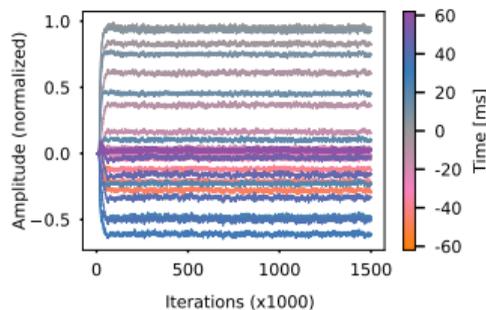
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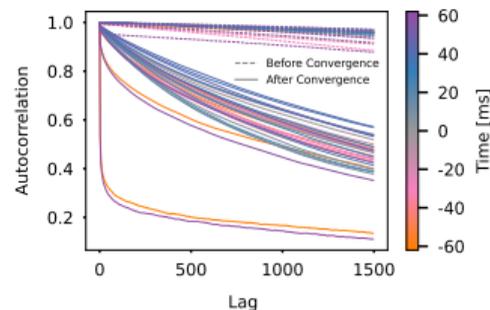
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## Marginal seismic wavelet estimation with Hamiltonian Monte Carlo

- The joint reflectivity and wavelet estimation is mathematically equivalent to sampling from the joint  $p(\mathbf{w}^*, \mathbf{c}^* | \mathbf{d}, \sigma_c^2, \sigma_w^2, \zeta)$ .
- We instead want a marginal wavelet estimation, i.e. to sample from  $p(\mathbf{w}^* | \mathbf{d}, \sigma_c^2, \sigma_w^2, \zeta)$ .
- This new marginal posterior looks like

$$p(\mathbf{w}^* | \mathbf{d}, \sigma_c^2, \sigma_w^2, \zeta, \mathbf{d}_{\text{aux}}) \propto p(\mathbf{w}^* | \sigma_w^2) p(\mathbf{d} | \mathbf{w}^*, \sigma_c^2, \sigma_w^2, \zeta, \mathbf{d}_{\text{aux}}),$$

with

$$p(\mathbf{d} | \mathbf{w}^*, \sigma_c^2, \sigma_w^2, \zeta, \mathbf{d}_{\text{aux}}) = \int p(\mathbf{d}, \mathbf{c}^* | \mathbf{w}^*, \sigma_c^2, \sigma_w^2, \zeta) d\mathbf{c}^* \quad (1)$$

- This integral can be done analytically.
- Using HMC might be better because it's a gradient-based algorithm.
- Evaluating gradients and densities at each iteration is  $\mathcal{O}(s^{1.5})$  on the cyclic lattice.

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- This integral can be done analytically.
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## Marginal seismic wavelet estimation with Hamiltonian Monte Carlo

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# Conclusions

- We proposed an approach for estimating a seismic wavelet with full UQ.
- Gibbs sampler allows joint wavelet and reflectivity estimation but is slow.
- Collapsed HMC allows efficient marginal wavelet estimation.

## References

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